

10-701: Introduction to Machine Learning

Lecture 7 - Naïve Bayes

Henry Chai

9/20/23

Front Matter

- Announcements:
 - HW1 released 9/6, due 9/20 (today!) at 11:59 PM
 - HW2 released 9/20 (today!), due 10/4 at 11:59 PM
 - Instructor OH have been added to the course calendar
 - Monday's immediately after lecture
- Recommended Readings:
 - Murphy, [Section 3.5](#)

Bayesian Linear Regression

$$p(y|x, \omega)$$

$$p(\omega)$$

- If we assume a linear model with additive Gaussian noise $y = \omega^T x + \epsilon$ where $\epsilon \sim N(0, \sigma^2)$ $\rightarrow y \sim N(\omega^T x, \sigma^2)$... and a **general** (zero-mean) Gaussian prior on the weights ...

$$\omega \sim N(\mathbf{0}, \Sigma)$$

then the distribution over \mathbf{y} is

$$\mathbf{y} \sim N(X\mathbf{0} + \mathbf{0} = \mathbf{0}, X\Sigma X^T + \sigma^2 I)$$

Bayesian Linear Regression

- If we assume a linear model with additive Gaussian noise
 $y = \boldsymbol{\omega}^T \mathbf{x} + \epsilon$ where $\epsilon \sim N(0, \sigma^2) \rightarrow y \sim N(\boldsymbol{\omega}^T \mathbf{x}, \sigma^2) \dots$
and a **general** (zero-mean) Gaussian prior on the weights ...
 $\boldsymbol{\omega} \sim N(\mathbf{0}, \Sigma)$

then the *joint* distribution over \mathbf{y} and $\boldsymbol{\omega}$ is

$$\begin{bmatrix} \mathbf{y} \\ \boldsymbol{\omega} \end{bmatrix} \sim N \left(\begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} X\Sigma X^T + \sigma^2 I & \Sigma X^T \\ X\Sigma & \Sigma \end{bmatrix} \right)$$

Bayesian Linear Regression

- If we assume a linear model with additive Gaussian noise

$$\mathbf{y} = \boldsymbol{\omega}^T \mathbf{x} + \epsilon \text{ where } \epsilon \sim N(0, \sigma^2) \rightarrow \mathbf{y} \sim N(\boldsymbol{\omega}^T \mathbf{x}, \sigma^2) \dots$$

and a **general** (zero-mean) Gaussian prior on the weights ...

$$\boldsymbol{\omega} \sim N(\mathbf{0}, \Sigma)$$

then the *conditional* distribution of $\boldsymbol{\omega}$ given \mathbf{y} is

$$\boldsymbol{\omega} \mid \mathbf{y} \sim N(\boldsymbol{\mu}_{POST}, \Sigma_{POST})$$

where

$$\boldsymbol{\mu}_{POST} = \Sigma X^T (X \Sigma X^T + \sigma^2 I)^{-1} \mathbf{y},$$

$$\Sigma_{POST} = \Sigma - \Sigma X^T (X \Sigma X^T + \sigma^2 I)^{-1} X \Sigma$$

Bayesian Linear Regression

- If we assume a linear model with additive Gaussian noise

$$y = \boldsymbol{\omega}^T \mathbf{x} + \epsilon \text{ where } \epsilon \sim N(0, \sigma^2) \rightarrow y \sim N(\boldsymbol{\omega}^T \mathbf{x}, \sigma^2) \dots$$

and a **general** (zero-mean) Gaussian prior on the weights ...

$$\boldsymbol{\omega} \sim N(\mathbf{0}, \Sigma)$$

then the *conditional* distribution of $h(\mathbf{x}') = \mathbf{x}'^T \boldsymbol{\omega}$ given \mathbf{y} is

$$h(\mathbf{x}') | \mathbf{y} \sim N(\boldsymbol{\mu}_{PRED}, \Sigma_{PRED})$$

where

$$\begin{aligned} \boldsymbol{\mu}_{PRED} &= \mathbf{x}'^T \Sigma X^T (X \Sigma X^T + \sigma^2 I)^{-1} \mathbf{y}, \\ \Sigma_{PRED} &= \mathbf{x}'^T \Sigma \mathbf{x}' - \mathbf{x}'^T \Sigma X^T (X \Sigma X^T + \sigma^2 I)^{-1} X \Sigma \mathbf{x}' \end{aligned}$$

Kernelized Bayesian Linear Regression = Gaussian Process (GP)

- If we assume a linear model with additive Gaussian noise

$$y = \boldsymbol{\omega}^T \mathbf{x} + \epsilon \text{ where } \epsilon \sim N(0, \sigma^2) \rightarrow y \sim N(\boldsymbol{\omega}^T \mathbf{x}, \sigma^2) \dots$$

and a **general** (zero-mean) Gaussian prior on the weights ...

$$\boldsymbol{\omega} \sim N(\mathbf{0}, \Sigma)$$

then the *conditional* distribution of $h(\mathbf{x}') = \mathbf{x}'^T \boldsymbol{\omega}$ given \mathbf{y} is

$$h(\mathbf{x}') | \mathbf{y} \sim N(\boldsymbol{\mu}_{PRED}, \Sigma_{PRED})$$

where

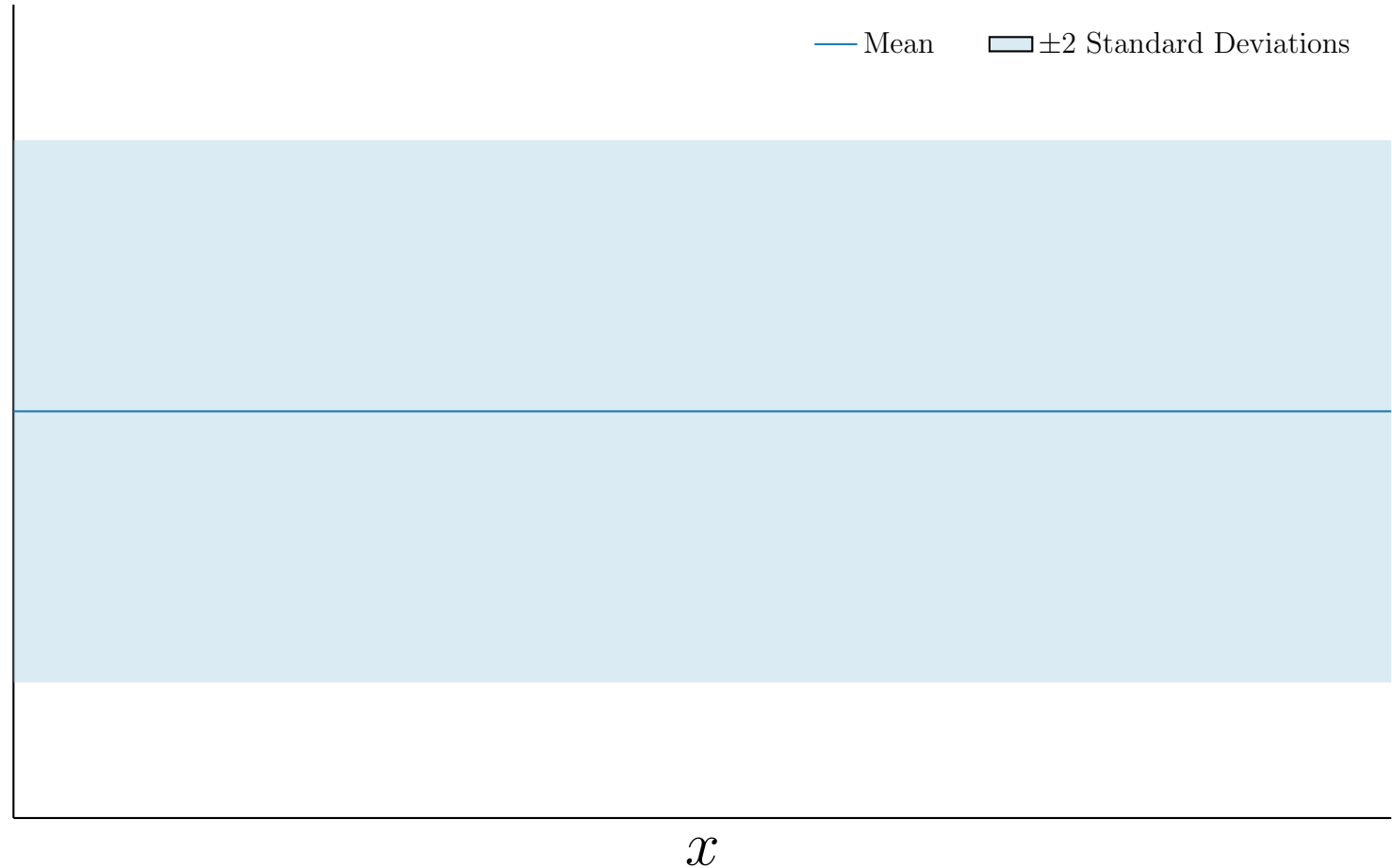
$$K(\mathbf{a}, \mathbf{b}) = \Phi(\mathbf{a})^T \Sigma \Phi(\mathbf{b})$$

$$\boldsymbol{\mu}_{PRED} = K(\mathbf{x}', X)(K(X, X) + \sigma^2 I)^{-1} \mathbf{y},$$

$$\Sigma_{PRED} = K(\mathbf{x}', \mathbf{x}') - K(\mathbf{x}', X)(K(X, X) + \sigma^2 I)^{-1} K(X, \mathbf{x}')$$

Gaussian Process (GP)

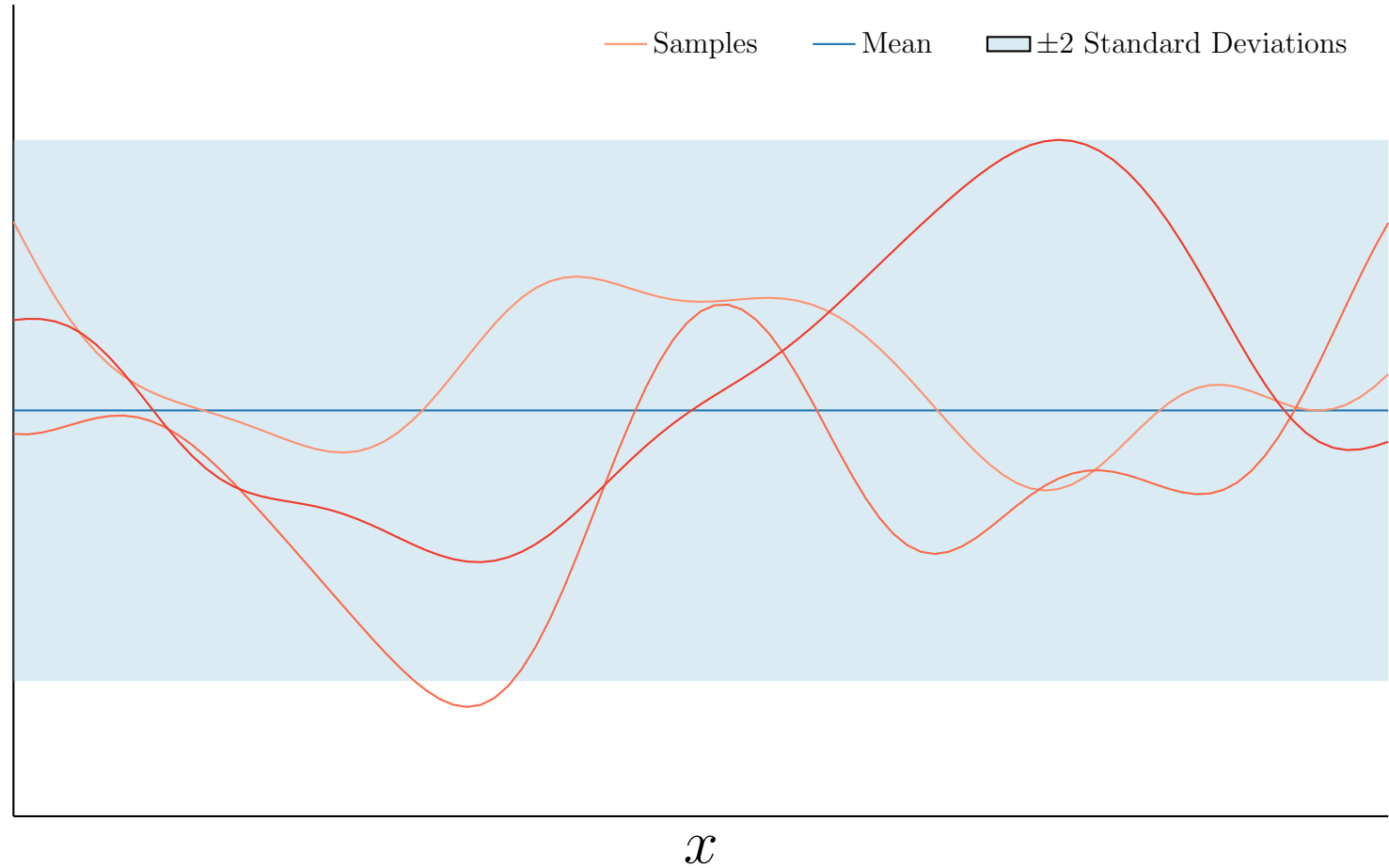
$$f \sim \mathcal{GP}(m(x) = 0, K(x, x') = \exp(-(x - x')^2))$$



$$f \sim \mathcal{GP}(m, K) \rightarrow f(x) \sim \mathcal{N}(m(x), K(x, x))$$

Gaussian Process (GP)

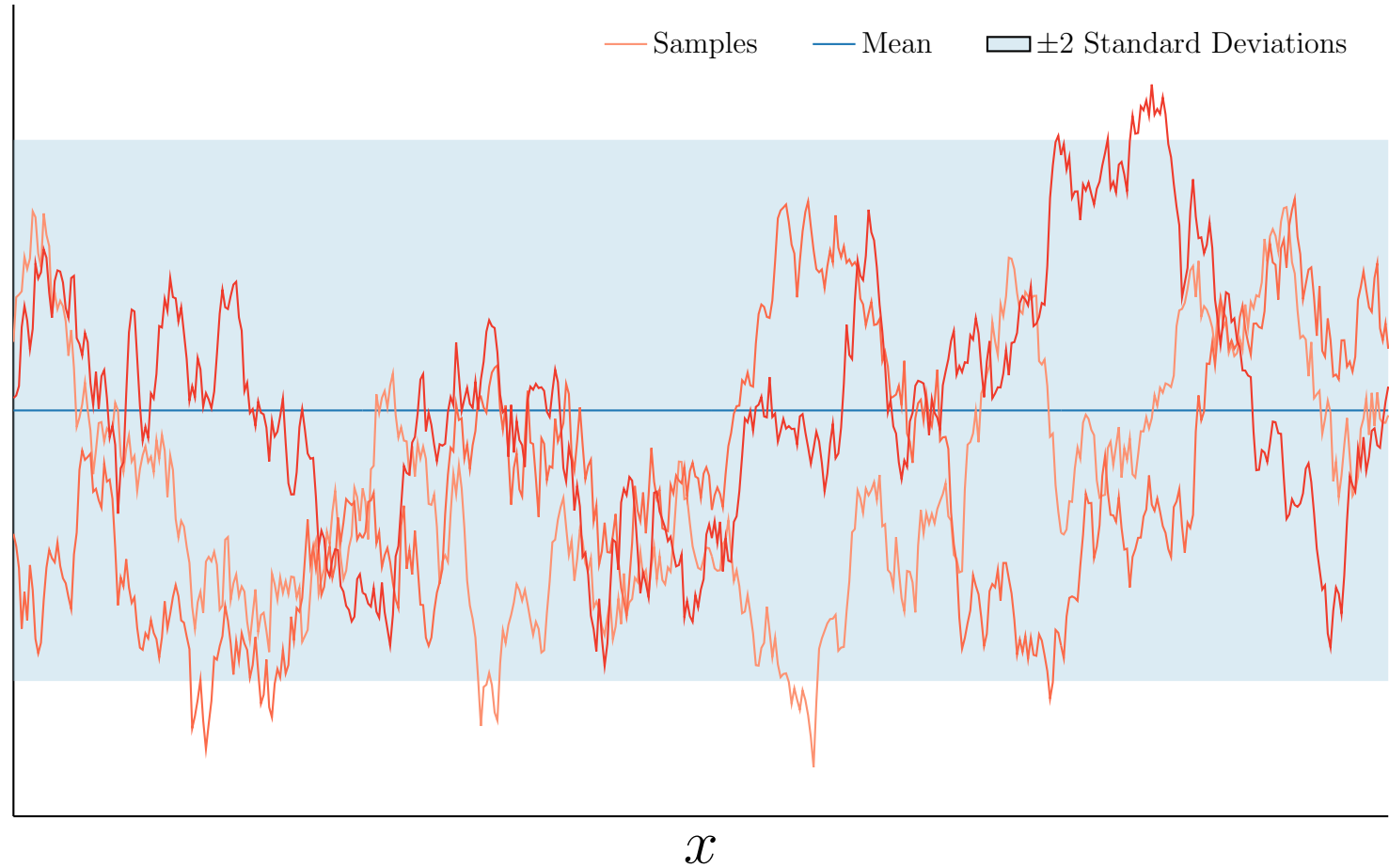
$$f \sim \mathcal{GP}(m(x) = 0, K(x, x') = \exp(-(x - x')^2))$$



$$f \sim \mathcal{GP}(m, K) \rightarrow f(x) \sim \mathcal{N}(m(x), K(x, x))$$

Gaussian Process (GP)

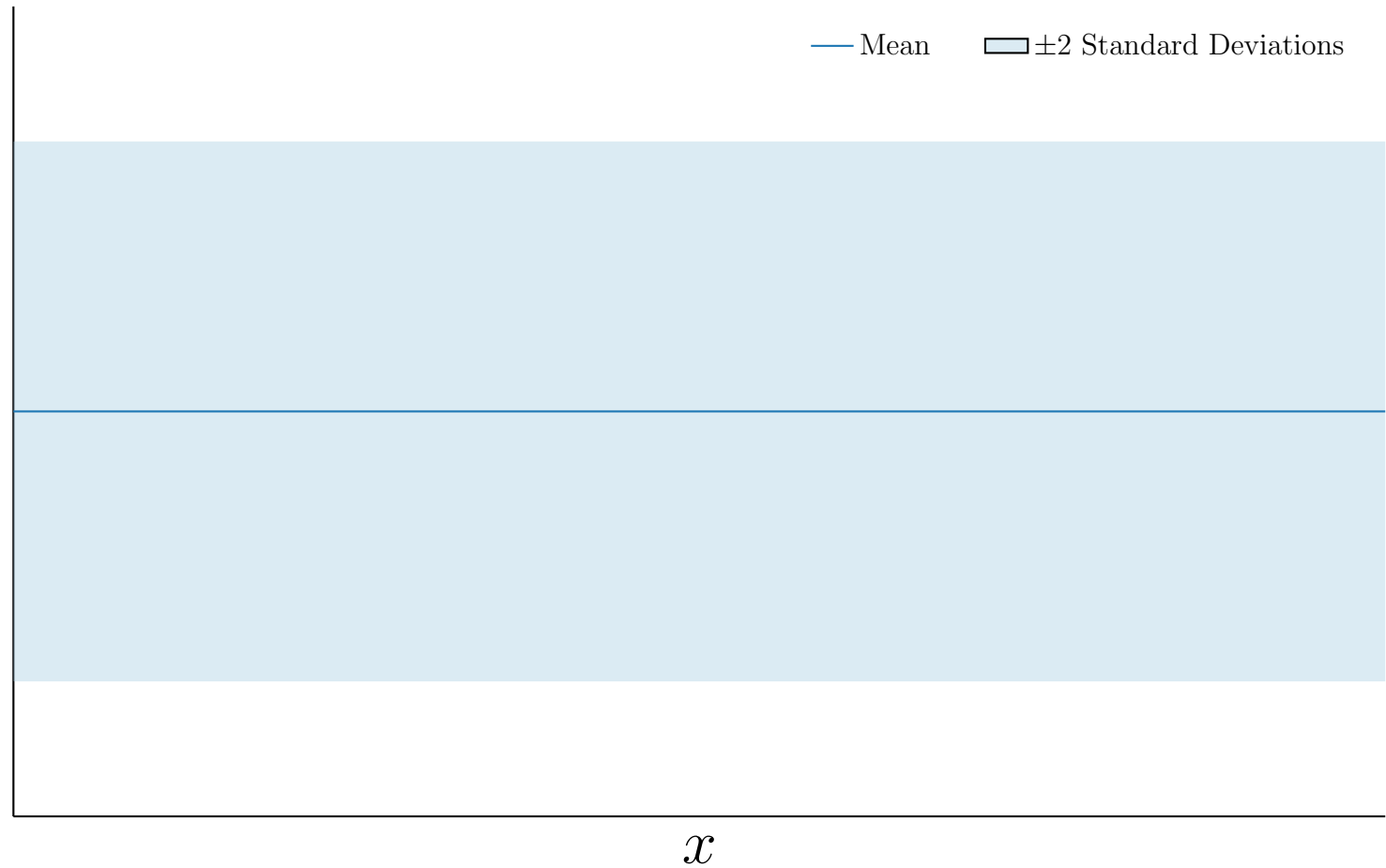
$$f \sim \mathcal{GP}(m(x) = 0, K(x, x') = \exp(-\|x - x'\|))$$



$$f \sim \mathcal{GP}(m, K) \rightarrow f(x) \sim \mathcal{N}(m(x), K(x, x))$$

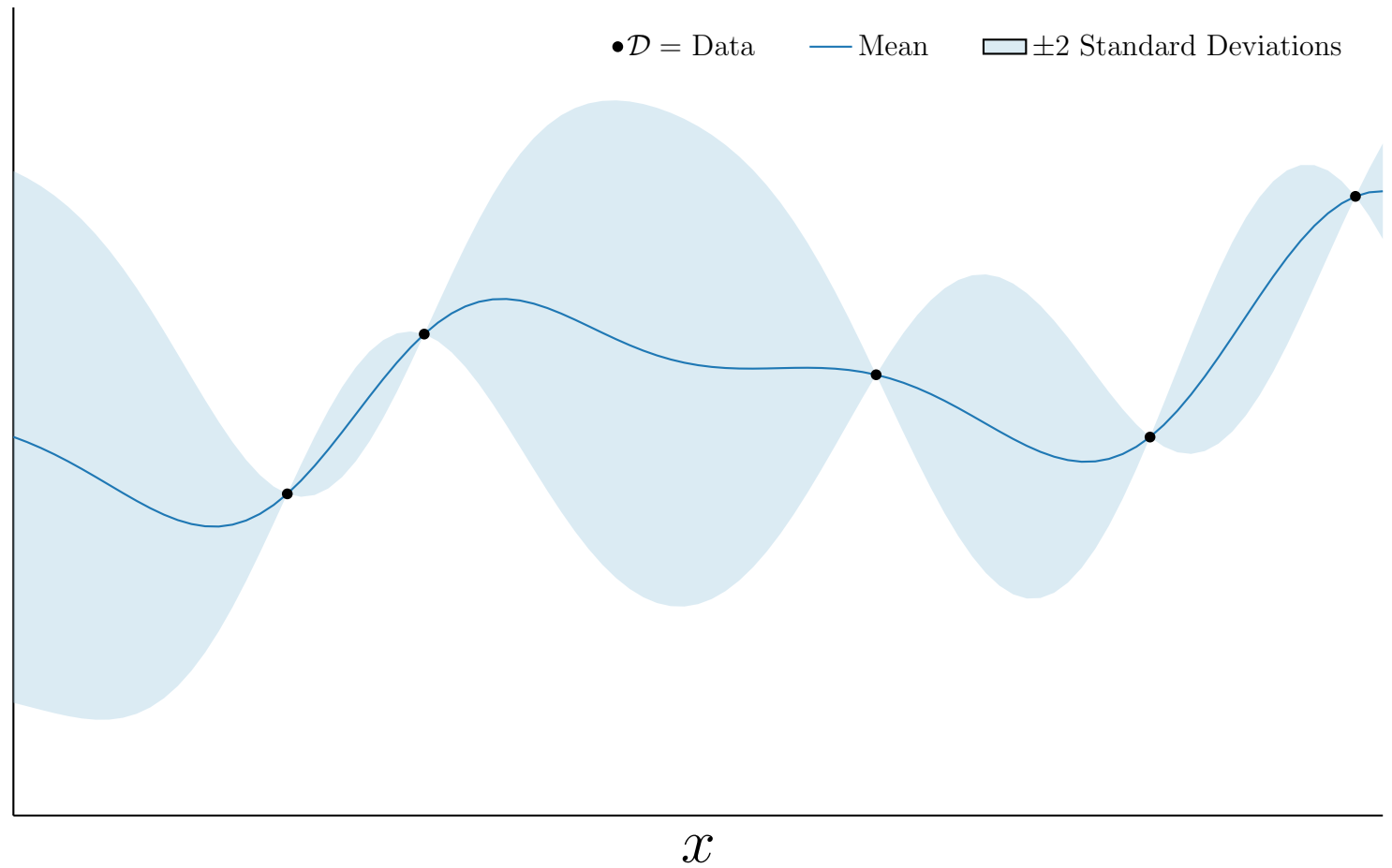
GP Prior

$$f \sim \mathcal{GP}(m(x) = 0, K(x, x') = \exp(-(x - x')^2))$$



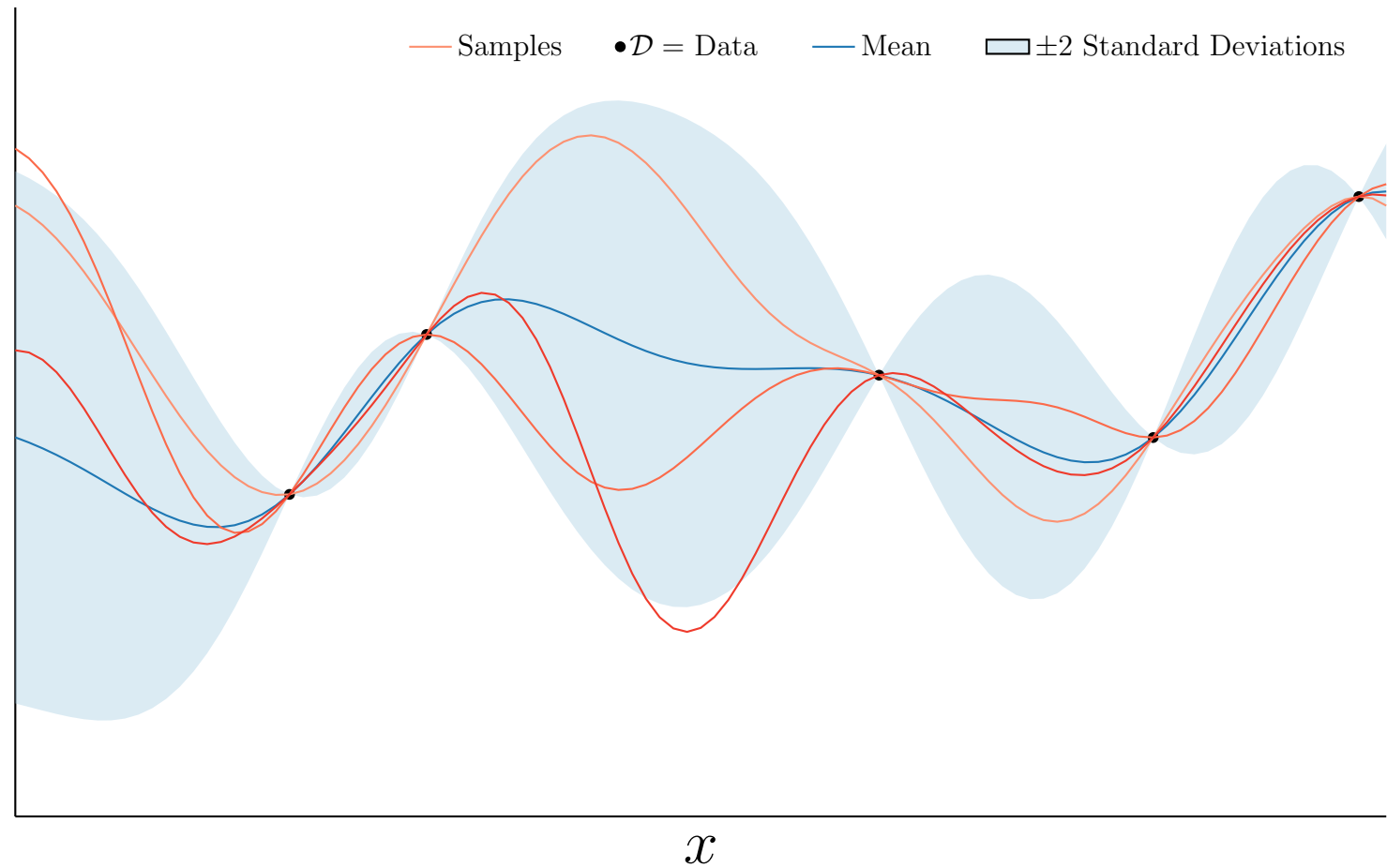
GP Posterior

$$f | \mathcal{D} \sim \mathcal{GP}(m_{\mathcal{D}}, K_{\mathcal{D}})$$



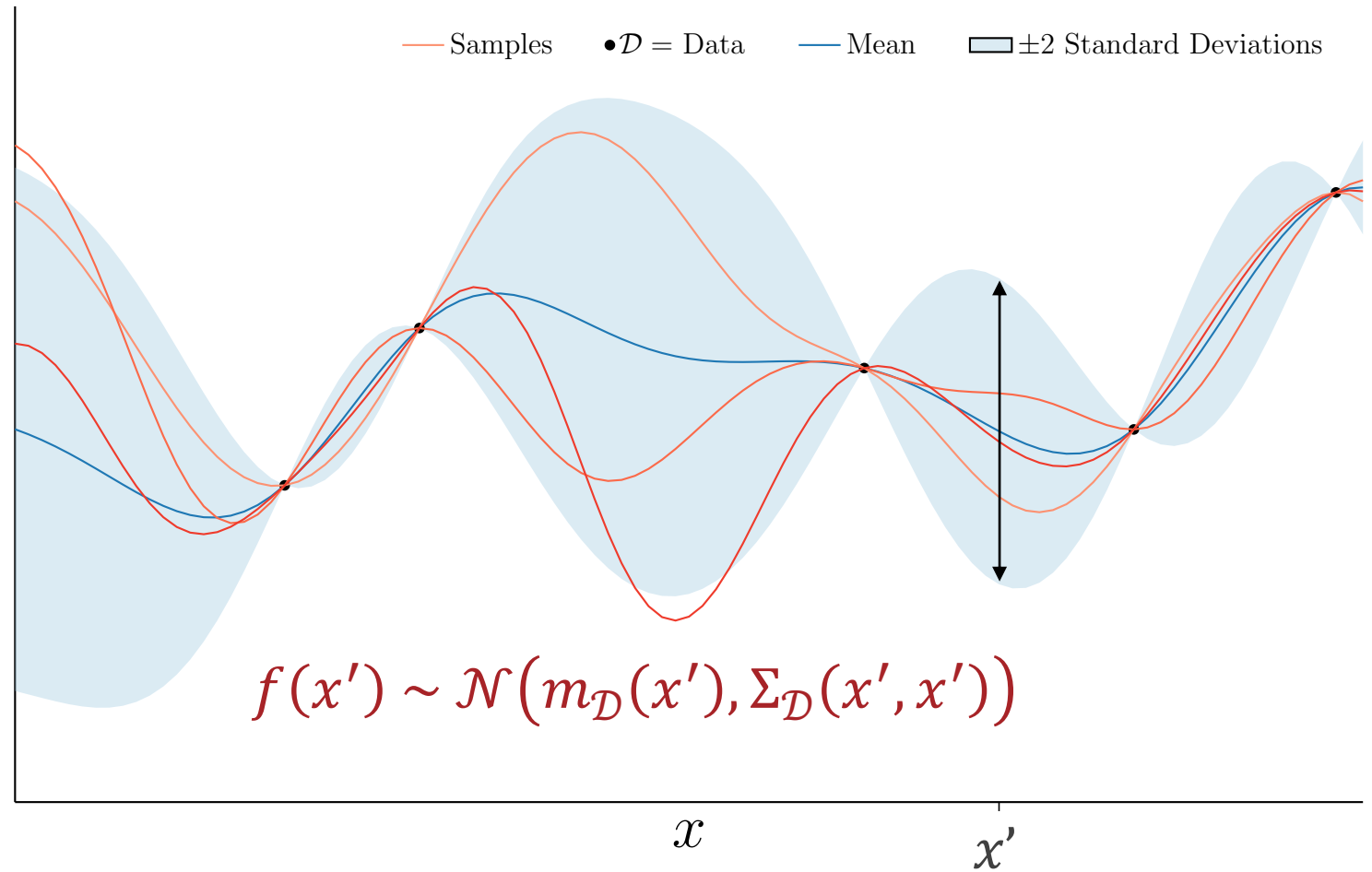
GP Posterior

$$f | \mathcal{D} \sim \mathcal{GP}(m_{\mathcal{D}}, K_{\mathcal{D}})$$



GP Posterior

$$f | \mathcal{D} \sim \mathcal{GP}(m_{\mathcal{D}}, K_{\mathcal{D}})$$



Text Data

- <https://www.nytimes.com/2023/09/19/us/politics/senate-dress-code-fetterman-schumer.html>
- <https://americanwirenews.com/slobs-of-the-world-unite-schumer-changes-senate-dress-code-to-accommodate-fetterman/>
- <https://triblive.com/news/pennsylvania/u-s-senate-loosens-dress-code-scoring-win-for-casually-dressed-fetterman/>
- <https://www.theonion.com/fetterman-struggling-to-adapt-to-size-of-capitol-buildi-1849773669>

Lawmakers Give New Senate Dress

Home PENNSYLVANIA

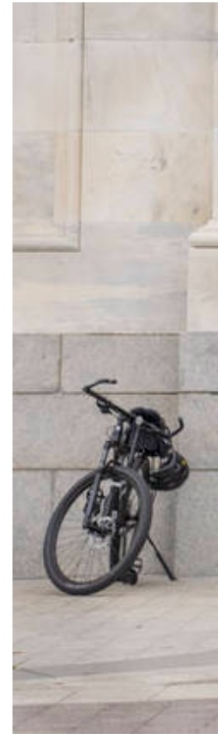
U.S. Sen
for casual

POLITICS

TRIB
LIVE RYAN DETO

Fetterman Struggling To Adapt To Size Of Capitol Building

Published January 4, 2023



AP

U.S. Sen. John Fetterman, D-Braddock, waves to members of the media on Capitol Hill in Washington on April 17, 2023.

Senate official. The

change applies only to senators — staff members will still be required to follow the old dress code,” Axios reported.

jes

tional
hat
rtinue

ness
ves

Text Data



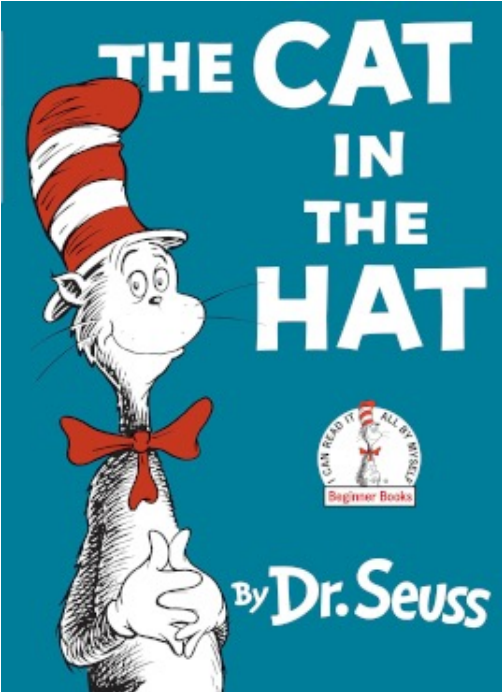
Bag-of-Words Model

x_1 ("hat")	x_2 ("cat")	x_3 ("dog")	x_4 ("fish")	x_5 ("mom")	x_6 ("dad")	y (Dr. Seuss)
------------------	------------------	------------------	-------------------	------------------	------------------	--------------------

Bag-of-Words Model

x_1 ("hat")	x_2 ("cat")	x_3 ("dog")	x_4 ("fish")	x_5 ("mom")	x_6 ("dad")	y (Dr. Seuss)
1	1	0	0	0	0	1

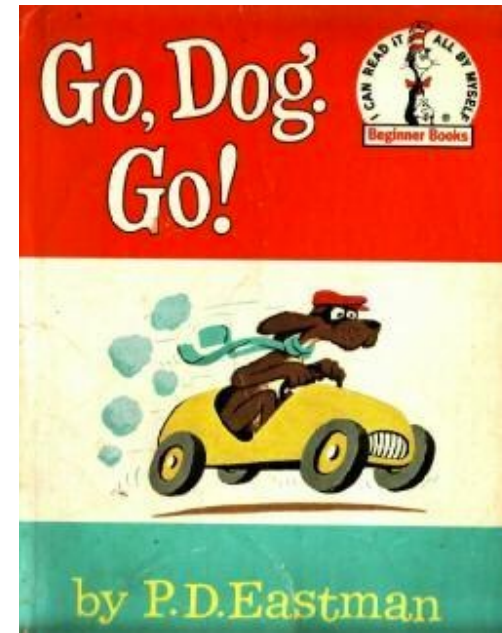
The **Cat** in the **Hat**
(by Dr. Seuss)



Bag-of-Words Model

x_1 ("hat")	x_2 ("cat")	x_3 ("dog")	x_4 ("fish")	x_5 ("mom")	x_6 ("dad")	y (Dr. Seuss)
1	1	0	0	0	0	1
0	0	1	0	0	0	0

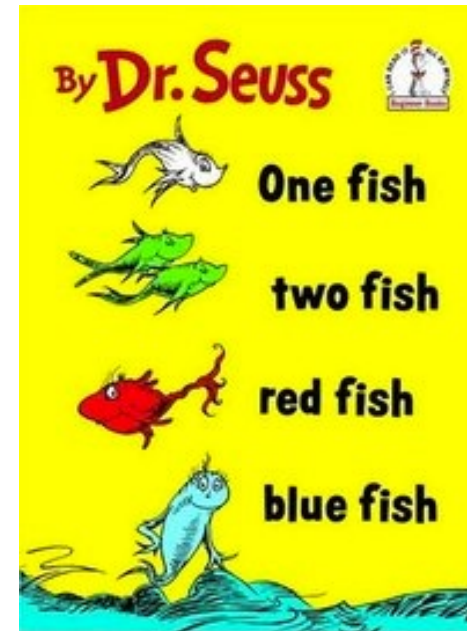
Go, **Dog**. Go!
(by P. D. Eastman)



Bag-of-Words Model

x_1 ("hat")	x_2 ("cat")	x_3 ("dog")	x_4 ("fish")	x_5 ("mom")	x_6 ("dad")	y (Dr. Seuss)
1	1	0	0	0	0	1
0	0	1	0	0	0	0
0	0	0	1	0	0	1

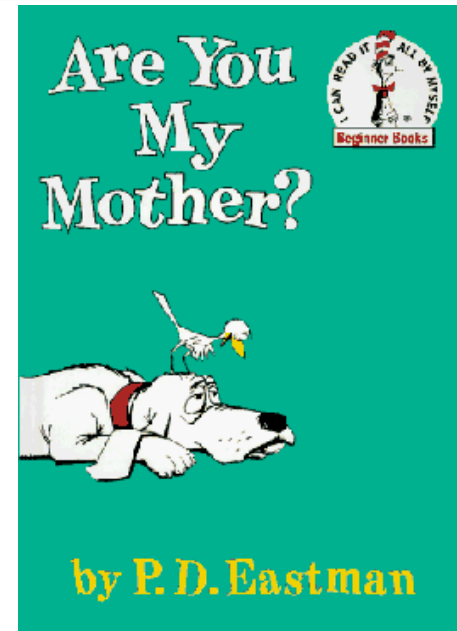
One Fish, Two Fish,
Red Fish, Blue Fish
(by Dr. Seuss)



Bag-of-Words Model

x_1 ("hat")	x_2 ("cat")	x_3 ("dog")	x_4 ("fish")	x_5 ("mom")	x_6 ("dad")	y (Dr. Seuss)
1	1	0	0	0	0	1
0	0	1	0	0	0	0
0	0	0	1	0	0	1
0	0	0	0	1	0	0

Are You My **Mother**?
(by P. D. Eastman)



Building a Probabilistic Classifier

- Define a decision rule
 - Given a test data point \mathbf{x}' , predict its label \hat{y} using the posterior distribution $P(Y = y|X = \mathbf{x}')$
 - Common choice: $\hat{y} = \underset{y}{\operatorname{argmax}} P(Y = y|X = \mathbf{x}')$
- Model the posterior distribution
 - Option 1 - Model $P(Y|X)$ directly as some function of X (later)
 - Option 2 - Use Bayes' rule (today!):

$$P(Y|X) = \frac{P(X|Y) P(Y)}{P(X)} \propto P(X|Y) P(Y)$$

likelihood

prior

How hard is modelling $P(X|Y)$?

- Define a decision rule
 - Given a test data point \mathbf{x}' , predict its label \hat{y} using the posterior distribution $P(Y = y|X = \mathbf{x}')$
 - Common choice: $\hat{y} = \underset{y}{\operatorname{argmax}} P(Y = y|X = \mathbf{x}')$
- Model the posterior distribution
 - Option 1 - Model $P(Y|X)$ directly as some function of X (later)
 - Option 2 - Use Bayes' rule (today!):

$$P(Y|X) = \frac{P(X|Y) P(Y)}{P(X)} \propto P(X|Y) P(Y)$$

How hard is modelling $P(X|Y)$?

$$P(x_i = 1 | Y=1) \Rightarrow P(x_i = 0 | Y=1) = 1 - P(x_i = 1 | Y=1)$$

$$P(x_i = 1 | Y=0)$$

↓ ↓

x_1 ("hat")	x_2 ("cat")	x_3 ("dog")	x_4 ("fish")	x_5 ("mom")	x_6 ("dad")	$P(X Y = 1)$
0	0	0	0	0	0	θ_1
1	0	0	0	0	0	θ_2
1	1	0	0	0	0	θ_3
1	0	1	0	0	0	θ_4

Naïve Bayes Assumption

- **Assume** features are conditionally independent given the label:

$$P(x|Y) = P(x_1 \cap x_2 \dots x_D | Y) = \prod_{d=1}^D P(x_d | Y)$$

- Pros:

Simplifies inference/computation (linear dependence on D)
Combats overfitting

- Cons:

It's a frequently incorrect assumption

General Recipe for Machine Learning

- Define a model and model parameters
- Write down an objective function
- Optimize the objective w.r.t. the model parameters

Recipe for Naïve Bayes

- Define a model and model parameters
 - Make the Naïve Bayes assumption
 - Assume iid data points
 - Parameters: $\pi = P(Y=1)$, $\Theta_{d,y} = P(x_d=1 | Y=y)$
- Write down an objective function
 - Maximize the log-likelihood
- Optimize the objective w.r.t. the model parameters
 - Solve for them in closed-form

Setting the Parameters via MLE

$$l_D(\pi, \theta) = \log P(\mathcal{D} = \{(x^{(i)}, y^{(i)})\}_{i=1}^N \mid \pi, \theta)$$

$$= \log \prod_{i=1}^N P(x^{(i)}, y^{(i)} \mid \pi, \theta)$$

$$= \sum_{i=1}^N \log (P(x^{(i)} \mid y^{(i)}, \theta) P(y^{(i)} \mid \pi))$$

$$= \sum_{i=1}^N \log (P(y^{(i)} \mid \pi) \prod_{d=1}^D P(x_d^{(i)} \mid y^{(i)}, \theta_{d,1}, \theta_{d,0}))$$

$$= \sum_{i=1}^N (\log(P(y^{(i)} \mid \pi))) + \sum_{d=1}^D \log (P(x_d^{(i)} \mid y^{(i)}, \theta_{d,1}, \theta_{d,0}))$$

$$= \sum_{i=1}^N (\log(\pi^{y^{(i)}} (1-\pi)^{1-y^{(i)}}))$$

$$+ \sum_{i: y^{(i)}=1} \sum_{d=1}^D \log(\theta_{d,1}^{x_d^{(i)}} (1-\theta_{d,1})^{1-x_d^{(i)}}) + \sum_{i: y^{(i)}=0} \dots$$

Setting the Parameters via MLE

- Binary label
 - $Y \sim \text{Bernoulli}(\pi)$
 - $\hat{\pi} = N_{Y=1} / N$
 - $N = \#$ of data points
 - $N_{Y=1} = \#$ of data points with label 1
- Binary features
 - $X_d | Y = y \sim \text{Bernoulli}(\theta_{d,y})$
 - $\hat{\theta}_{d,y} = N_{Y=y, X_d=1} / N_{Y=y}$
 - $N_{Y=y} = \#$ of data points with label y
 - $N_{Y=y, X_d=1} = \#$ of data points with label y and feature $X_d = 1$

Bernoulli Naïve Bayes

- Binary label
 - $Y \sim \text{Bernoulli}(\pi)$
 - $\hat{\pi} = N_{Y=1} / N$
 - $N = \#$ of data points
 - $N_{Y=1} = \#$ of data points with label 1
- Binary features
 - $X_d | Y = y \sim \text{Bernoulli}(\theta_{d,y})$
 - $\hat{\theta}_{d,y} = N_{Y=y, X_d=1} / N_{Y=y}$
 - $N_{Y=y} = \#$ of data points with label y
 - $N_{Y=y, X_d=1} = \#$ of data points with label y and feature $X_d = 1$

Multiclass Bernoulli Naïve Bayes

- Discrete label (Y can take on one of M possible values)
 - $Y \sim \text{Categorical}(\pi_1, \dots, \pi_M)$
 - $\hat{\pi}_m = N_{Y=m} / N$
 - $N = \#$ of data points
 - $N_{Y=m} = \#$ of data points with label m
- Binary features
 - $X_d | Y = m \sim \text{Bernoulli}(\theta_{d,m})$
 - $\hat{\theta}_{d,m} = N_{Y=m, X_d=1} / N_{Y=m}$
 - $N_{Y=m} = \#$ of data points with label m
 - $N_{Y=m, X_d=1} = \#$ of data points with label m and feature $X_d = 1$

Multinomial Naïve Bayes

- Binary label
 - $Y \sim \text{Bernoulli}(\pi)$
 - $\hat{\pi} = N_{Y=1} / N$
 - $N = \#$ of data points
 - $N_{Y=1} = \#$ of data points with label 1
- Discrete features (X_d can take on one of K possible values)
 - $X_d | Y = y \sim \text{Categorical}(\theta_{d,1,y}, \dots, \theta_{d,K,y})$
 - $\hat{\theta}_{d,k,y} = N_{Y=y, X_d=k} / N_{Y=y}$
 - $N_{Y=y} = \#$ of data points with label y
 - $N_{Y=y, X_d=k} = \#$ of data points with label y and feature $X_d = k$

Gaussian Naïve Bayes

- Binary label
 - $Y \sim \text{Bernoulli}(\pi)$
 - $\hat{\pi} = N_{Y=1}/N$
 - $N = \#$ of data points
 - $N_{Y=1} = \#$ of data points with label 1
- Real-valued features
 - $X_d | Y = y \sim \text{Gaussian}(\mu_{d,y}, \sigma_{d,y}^2)$
 - $\hat{\mu}_{d,y} = \frac{1}{N_{Y=y}} \sum_{n:y^{(n)}=y} x_d^{(n)}$
 - $\hat{\sigma}_{d,y}^2 = \frac{1}{N_{Y=y}} \sum_{n:y^{(n)}=y} (x_d^{(n)} - \hat{\mu}_{d,y})^2$
 - $N_{Y=y} = \#$ of data points with label y

Recall: Fisher Iris Dataset

- Fisher (1936) used 150 measurements of flowers from 3 different species: Iris setosa (0), Iris virginica (1), Iris versicolor (2) collected by Anderson (1936)

Species	Sepal Length	Sepal Width
0	4.3	3.0
0	4.9	3.6
0	5.3	3.7
1	4.9	2.4
1	5.7	2.8
1	6.3	3.3
1	6.7	3.0

Visualizing Gaussian Naïve Bayes (2 classes)

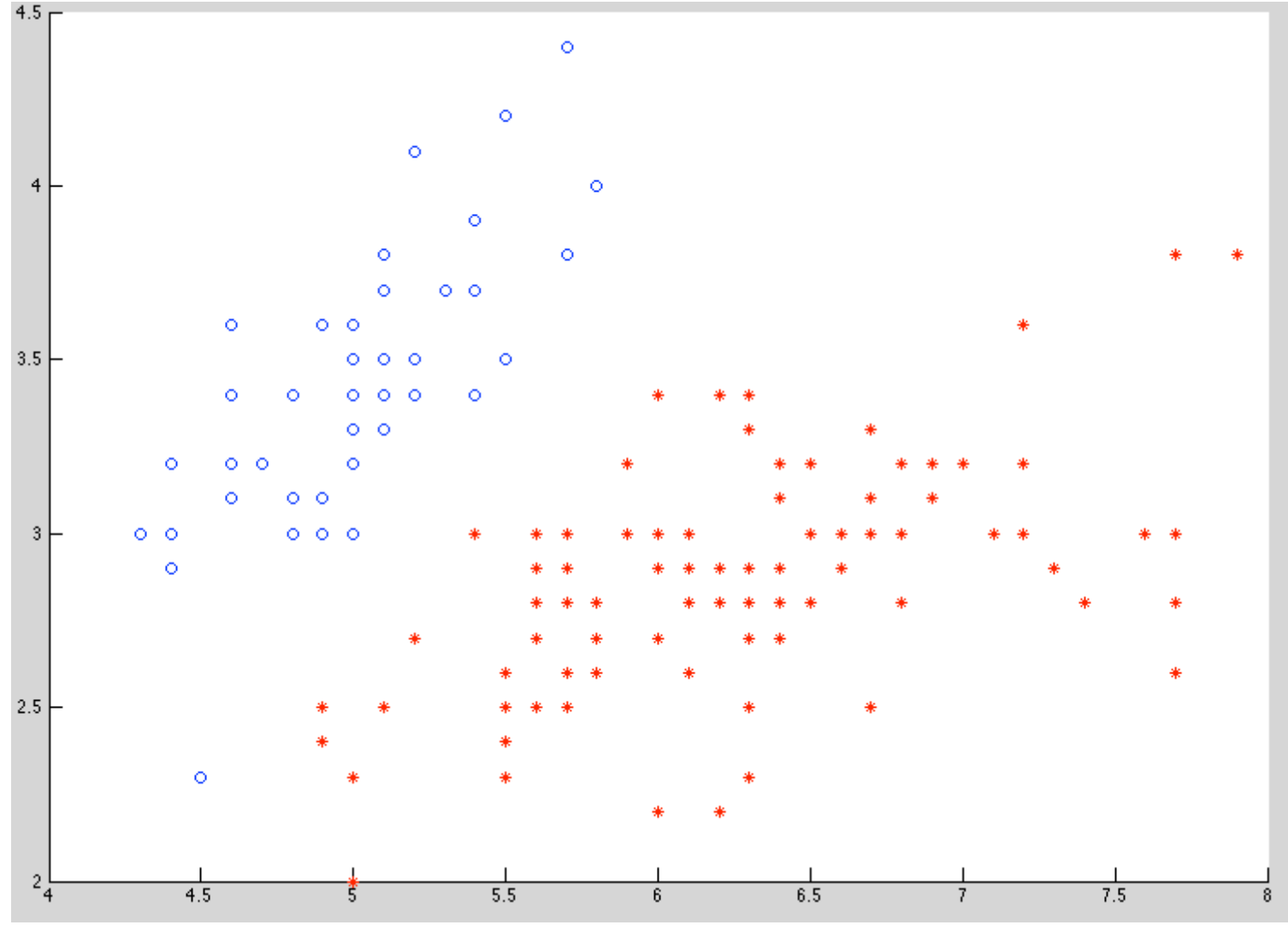
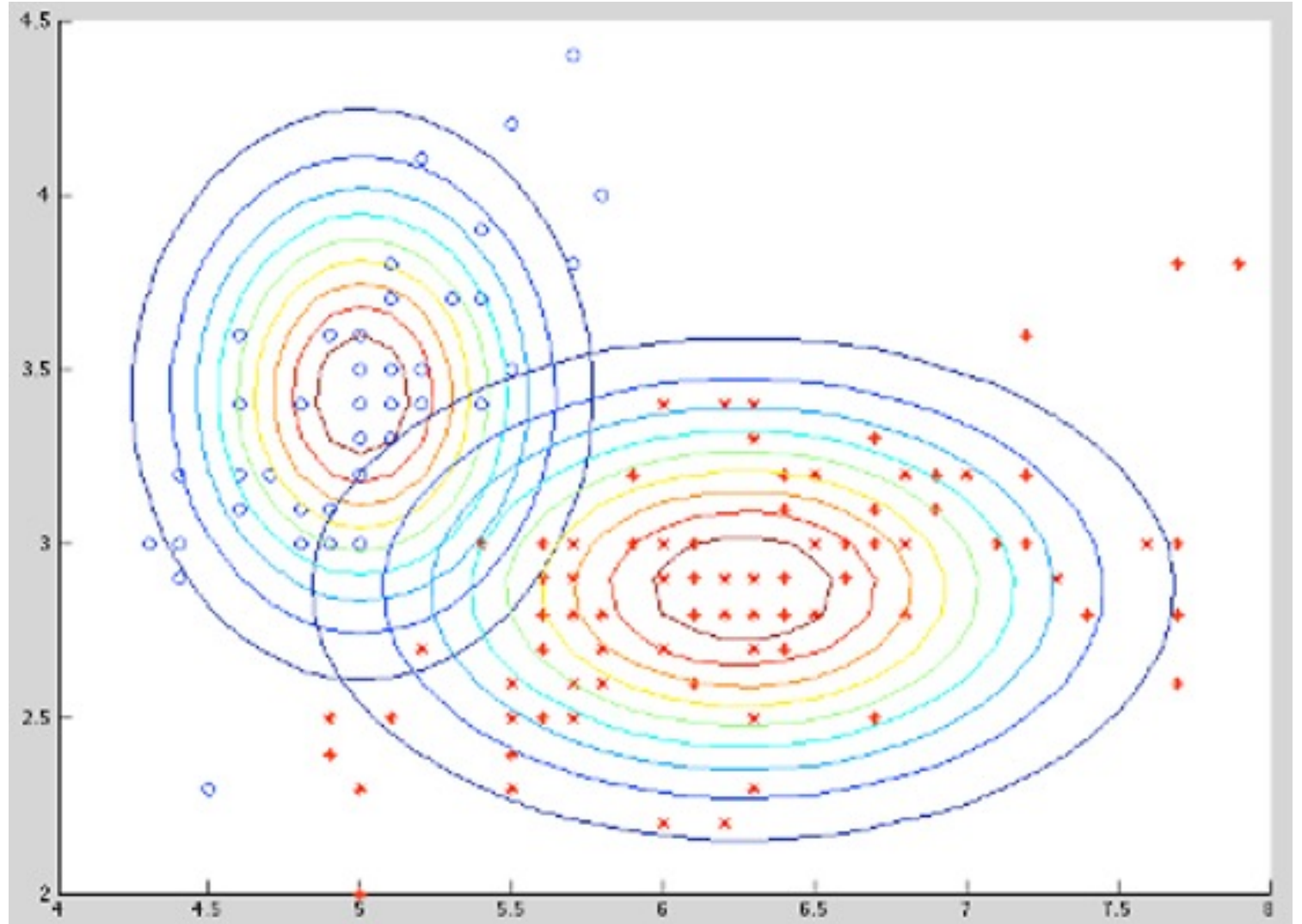


Figure courtesy of William Cohen

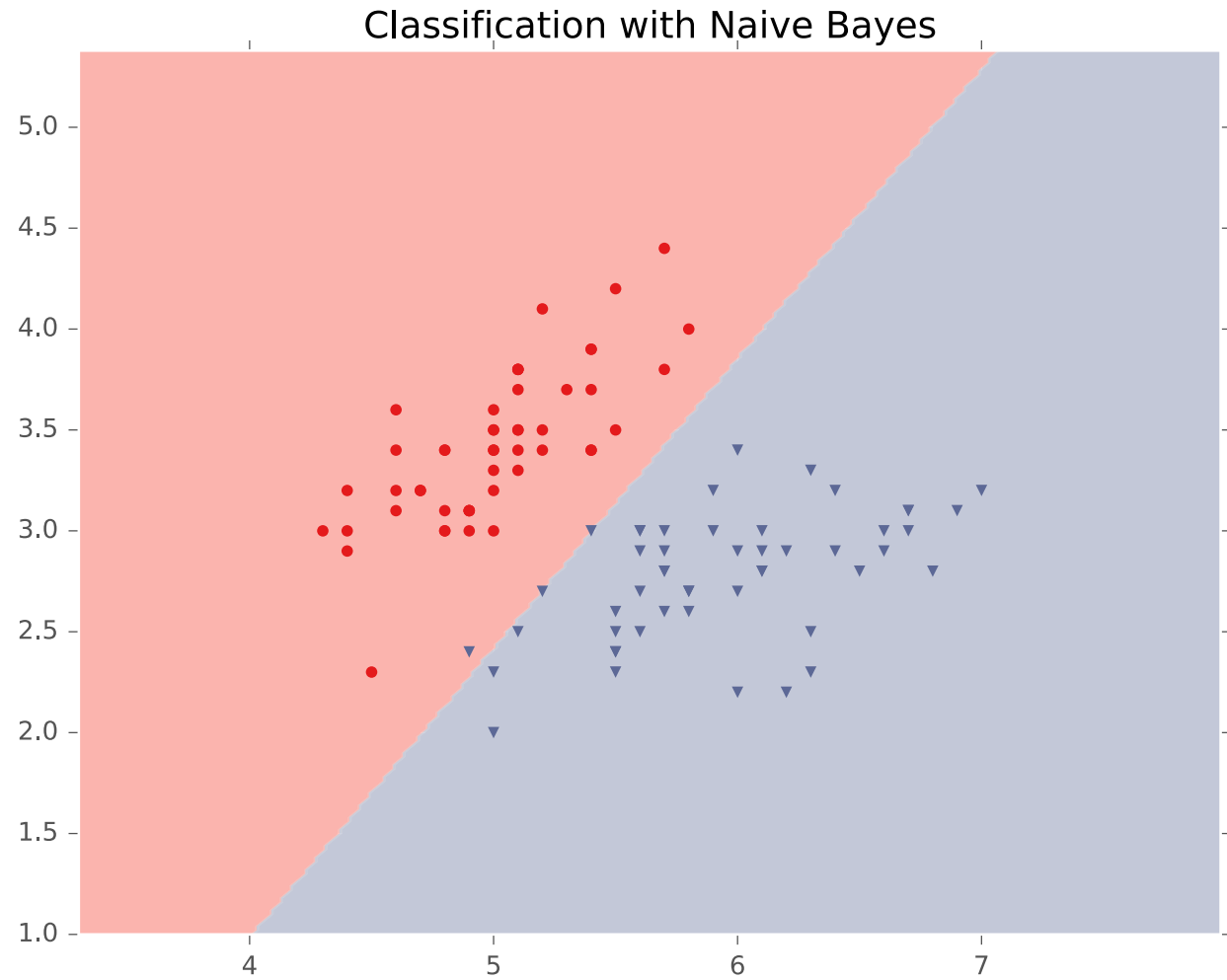
Visualizing Gaussian Naïve Bayes (2 classes)



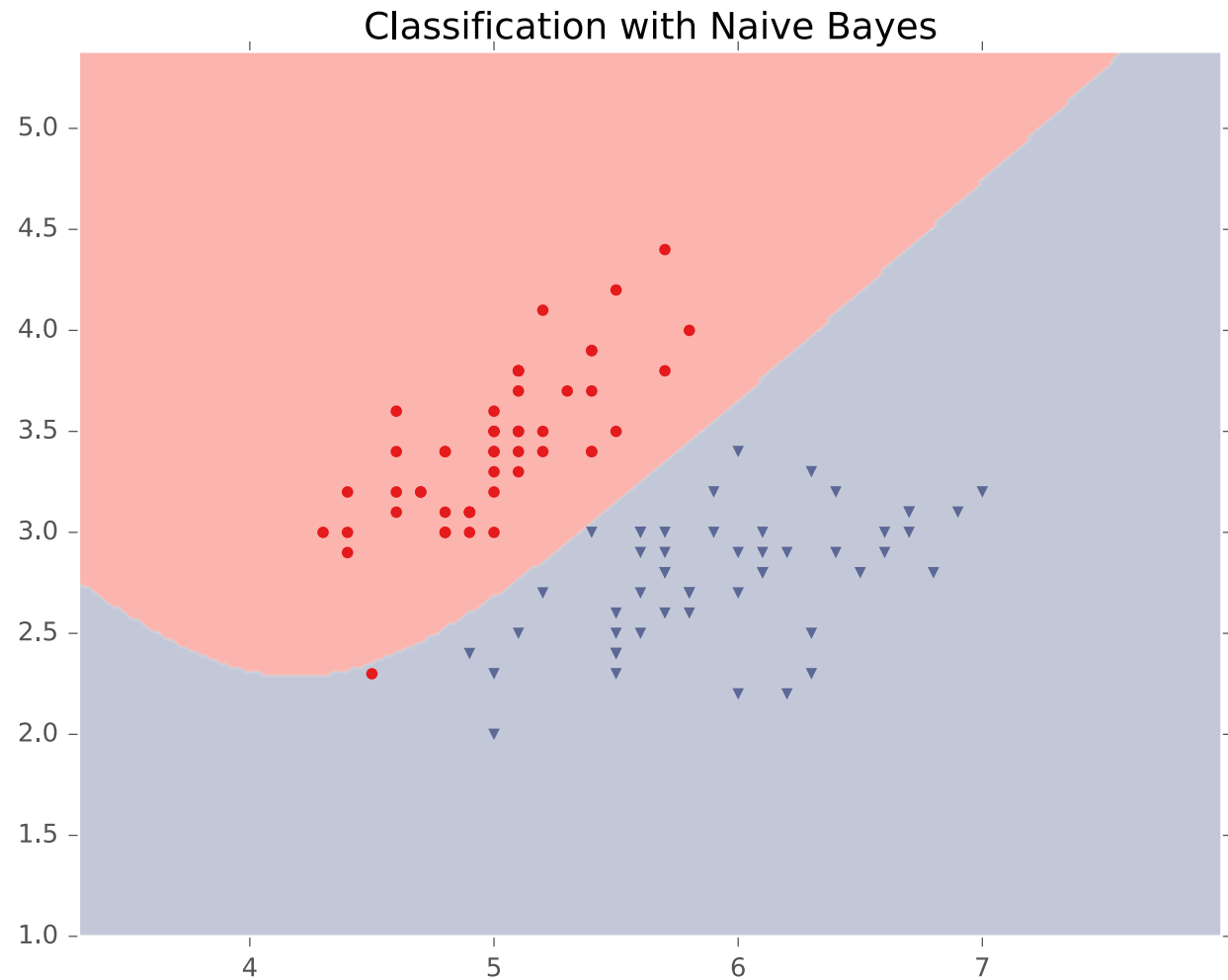
$$\hat{\sigma}_{d_1}^2 = \hat{\sigma}_{d_1}^2$$

$$\hat{\sigma}_{x_{1|1}}^2 = \hat{\sigma}_{x_{2|1}}^2$$

Visualizing
Gaussian
Naïve
Bayes
(2 classes,
equal
variances)



Visualizing Gaussian Naïve Bayes (2 classes, learned variances)



What if some
 Bernoulli
 variables
 appears in our
 training data?
 Predictions

- Given a test data point $\mathbf{x}' = [x'_1, \dots, x'_D]^T$

$$\underline{P(Y=1 | \mathbf{x}')} \propto P(\mathbf{x}' | Y=1) P(Y=1)$$

$$= \frac{\pi}{\pi} \prod_{d=1}^D \hat{\theta}_{d,1}^{x'_d} (1 - \hat{\theta}_{d,1})^{1-x'_d}$$

$$P(Y=0 | \mathbf{x}') \propto (1 - \frac{\pi}{\pi}) \prod_{d=1}^D \hat{\theta}_{d,0}^{x'_d} (1 - \hat{\theta}_{d,0})^{1-x'_d}$$

$$y = \begin{cases} 1 & \text{if } P(Y=1 | \mathbf{x}') \geq P(Y=0 | \mathbf{x}') \\ 0 & \text{otherwise} \end{cases}$$

What if some Word-Label pair never appears in our training data?

x_1 ("hat")	x_2 ("cat")	x_3 ("dog")	x_4 ("fish")	x_5 ("mom")	x_6 ("dad")	y (Dr. Seuss)
1	1	0	0	0	0	1
0	0	1	0	0	0	0
0	0	0	1	0	0	1
0	0	0	0	1	0	0

The Cat in the Hat gets a Dog (by ???)

- If some $\hat{\theta}_{d,y} = 0$ and that word appears in our test data \mathbf{x}' , then $P(Y = y|\mathbf{x}') = 0$ even if all the other features in \mathbf{x}' point to the label being y !
- The model has been overfit to the training data...
- We can address this with a prior over the parameters!

Setting the Parameters via MAP

- Binary label
 - $Y \sim \text{Bernoulli}(\pi)$
 - $\hat{\pi} = N_{Y=1} / N$
 - $N = \#$ of data points
 - $N_{Y=1} = \#$ of data points with label 1
- Binary features
 - $X_d | Y = y \sim \text{Bernoulli}(\theta_{d,y})$ and $\theta_{d,y} \sim \text{Beta}(\alpha, \beta)$
 - $\hat{\theta}_{d,y} = \frac{N_{Y=y, X_d=1} + (\alpha - 1)}{N_{Y=y} + (\alpha - 1) + (\beta - 1)}$
 - $N_{Y=y} = \#$ of data points with label y
 - $N_{Y=y, X_d=1} = \#$ of data points with label y and feature $X_d = 1$
 - α and β are “pseudocounts” of imagined data points that help avoid zero-probability predictions.
 - Common choice: $\alpha = \beta = 2$

What can we do when this is a bad/incorrect assumption, e.g., when our features are words in a sentence?

- **Assume** features are conditionally independent given the label:

$$P(X|Y) = \prod_{d=1}^D P(X_d|Y)$$

- Pros:
 - Significantly reduces computational complexity
 - Also reduces model complexity, combats overfitting
- Cons:
 - Is a strong, often illogical assumption
 - We'll see a relaxed version of this next week when we discuss Bayesian networks

Key Takeaways

- Text data
 - Bag-of-words feature representation
- Naïve Bayes
 - Conditional independence assumption
 - Pros and cons
 - Different Naïve Bayes models based on type of features
 - MLE vs. MAP for Bernoulli Naïve Bayes